

Multiple Questions for Multiple Mediators

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Mediation-type enquiries



Increasingly, mediation-type questions are posed in several contexts:

▶ RCTs:

Is cognitive behavioral therapy (CBT) acting via increased compliance to medication in reducing suicide rates?

▶ Public health interventions:

Does an enhanced school environment improve students' physical activity via their empowerment?

▶ Aetiological studies:

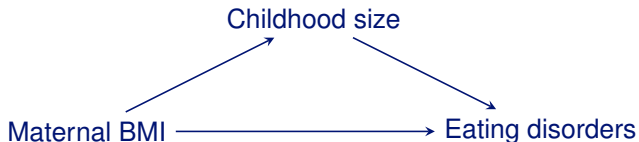
What proportion of the effect of prenatal factors on eating disorders can be explained by childhood growth?

Mediation



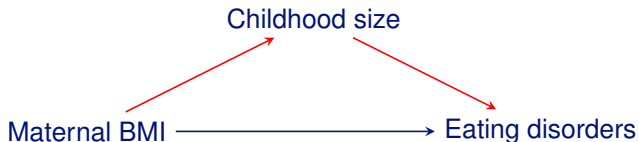
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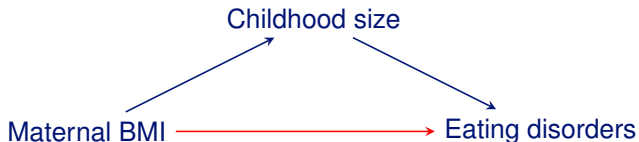
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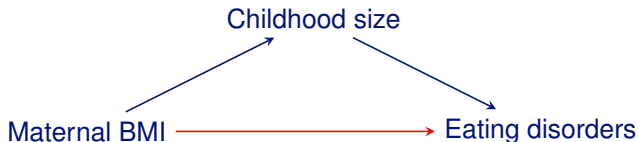
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Talk is about the complexities of attempting mediation analysis

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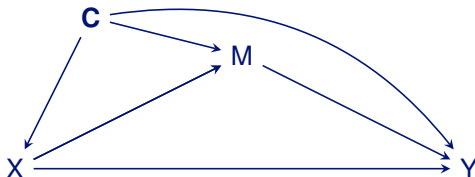
A long tradition



- Mediation analysis is not new; statisticians have been attempting it using traditional regression models for nearly a century
- It is inspired by **path analysis** (Wright, 1921, 1934)
- It is popular in the social sciences, while in epidemiology a less structured approach is often adopted
- In each of these approaches a **set of regression models** is specified

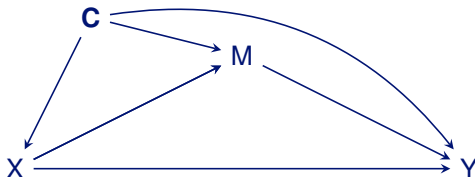
Path analysis for a simple setting

- Consider a structure with exposure X , continuous mediator M , continuous outcome Y and confounders C
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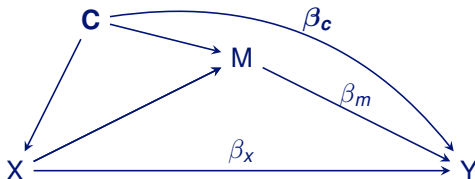
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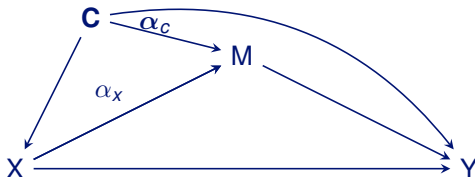
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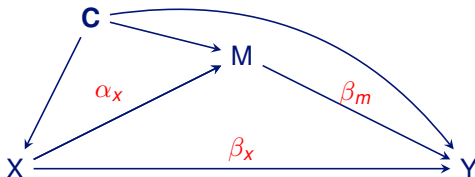


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Interpretation: — β_x as the **direct effect** of X on Y
 — $\alpha_x \beta_m$ as the **indirect effect** of X on Y (*path tracing rule*)

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- Path tracing rules are derived from **partitioning of correlations** (Wright, 1934): not based on a formal causal argument
- Yet, mediation enquiries make sense only if we think of them as causal questions:
 - When we say M mediates the effect of X on Y , we mean that X has a causal effect on Y and that some of this effect is explained by the fact that X has a causal effect on M which in turn has a causal effect on Y .

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2. Reliance on linearity



- If the model for Y were specified as:

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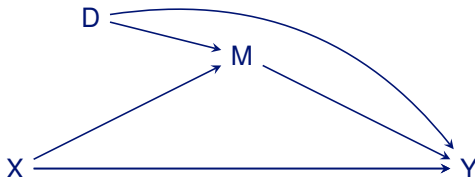
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- If Y were a binary variable modelled by **logistic regression**, even in the absence of an X – M interaction, application of the path tracing rule would not be sensible.

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3. M - Y confounding

It is not always sufficiently emphasised that common causes of M and Y should be controlled for. Consider this simple diagram:

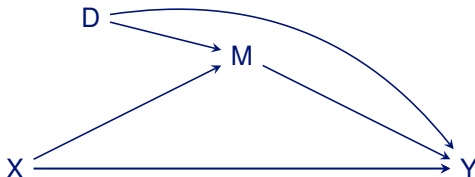


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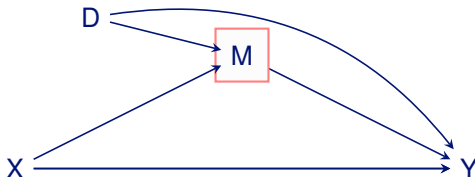
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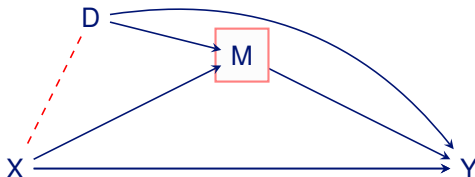
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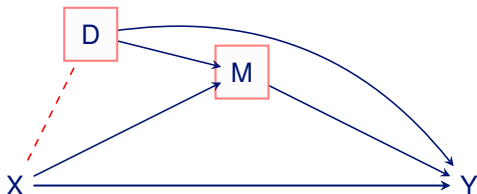


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- Conditioning on D too, as in:

$$E(Y) = \beta_0^* + \beta_x^* X + \beta_m^* M + \beta_d^* D$$

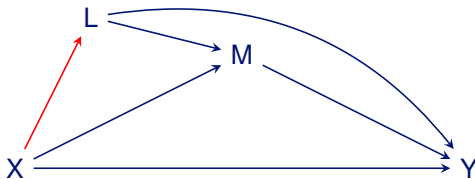
renders both paths blocked, so that β_x^* represents the direct effect.

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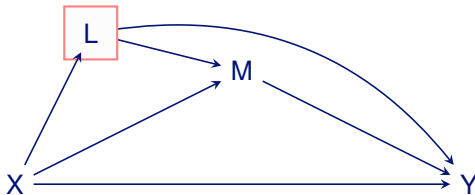


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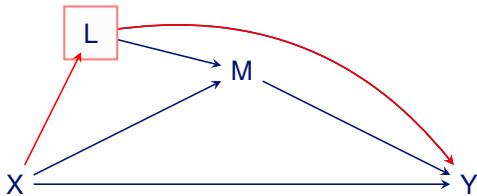
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- We have just seen that common causes of M and Y should be controlled for
- However this would block part of the direct effect from X to Y (via L but not via M)

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Counterfactual-based mediation



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- Causal inference involves the notion of *“How the world would have been had something been different”*
- It leads to defining *potential* (counterfactual) *outcomes* and *mediators*
- These counterfactuals are central to the definitions of direct/indirect effects in this framework

Counterfactuals



- $Y(x)$

the value that Y would take if we **intervened** on X and set it (possibly counter to fact) to the value x .

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- $Y(x, m)$

the value that Y would take if we intervened simultaneously on both X and M and set them to the values x and m .

$Y\{x, M(x^*)\}$

the value that Y would take if we intervened on X and set it to x whilst simultaneously intervening on M and setting it to $M(x^*)$, where x and x^* are not necessarily equal (Nested counterfactual).

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Mediation effects (estimands) are defined in terms of these potential entities, without invoking any parametric model

Natural direct effect

Pearl, 2001; Robins and Greenland, 1992



The average **natural direct effect** of X on Y expressed as a marginal mean difference is

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- This is a comparison of two hypothetical worlds.
- In the first, X is set to 1, and in the second X is set to 0. In **both** worlds, M is set to $M(0)$, the value it would take if X were set to 0.
- Since M is the same (*within* subject) in both worlds, we are still getting at the **direct effect** of X .

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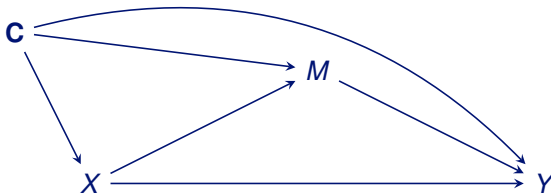
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The direct and indirect effects together sum to the total causal effect:

$$\text{TCE} = \text{NDE} + \text{NIE}$$

Identification of natural effects

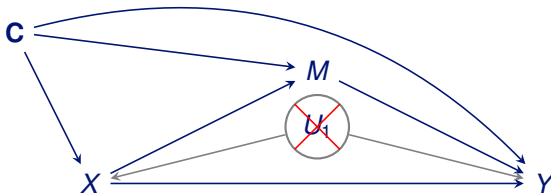
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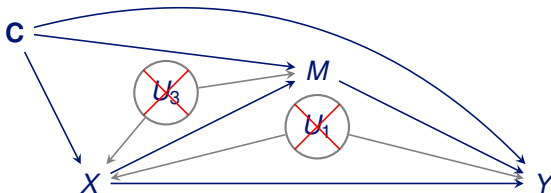
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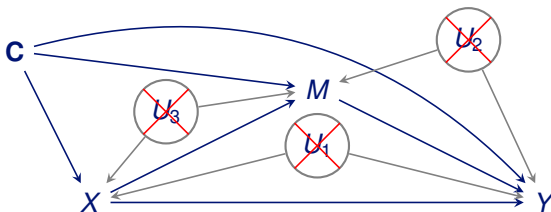
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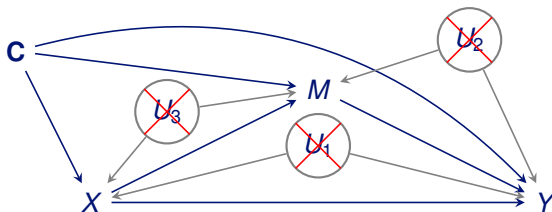
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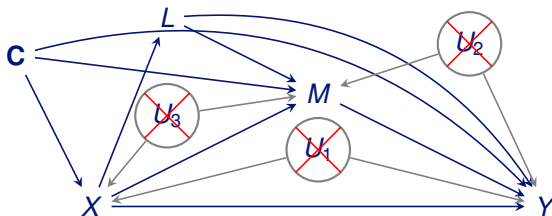
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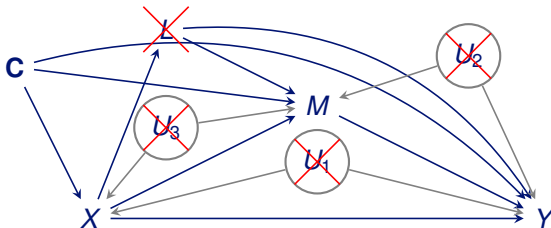
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- ▶ No need for the cross-world assumption for their identification; however, their sum is not TCE

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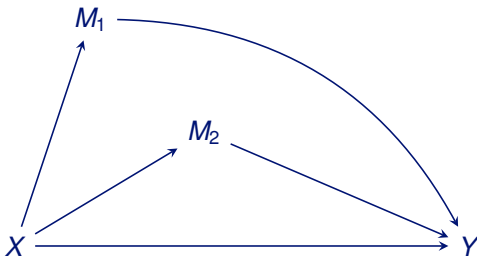


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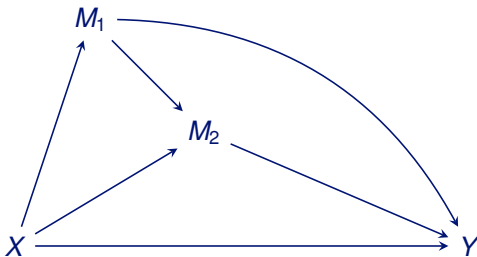
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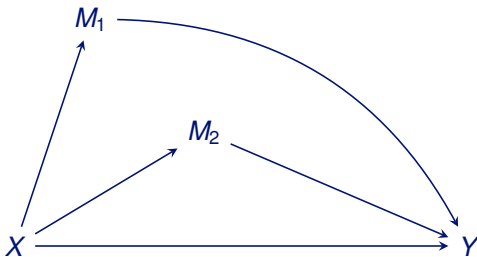
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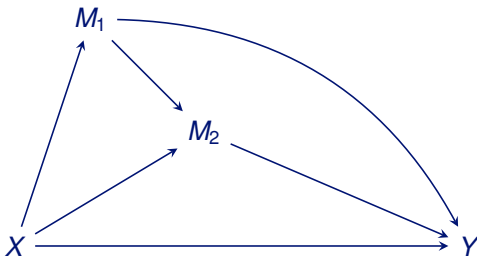
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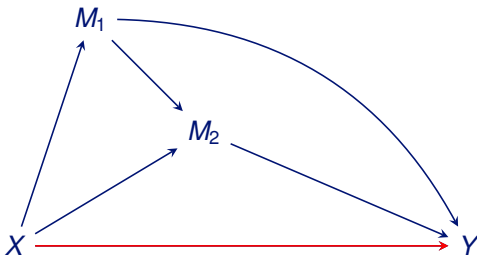


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4. If the mediators are **repeated measures** of the same dimension, focus on interventional effects of time-varying mediators [Lin *et al.* 2016, VanderWeele & Tchetgen-Tchetgen 2017]

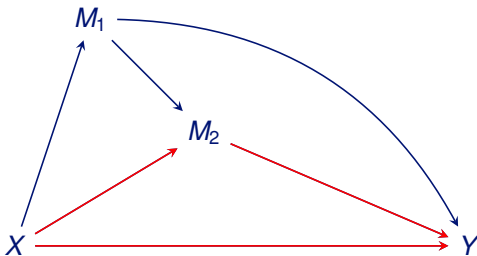
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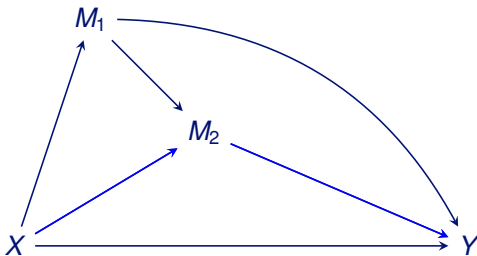
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1. & 2. Multivariable and sequential mediation

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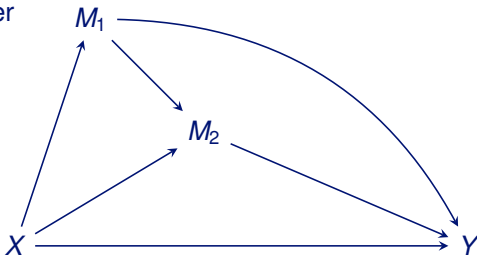


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These steps lead to partition of TCE into 3 components:

- via M_1 and downstream
- via M_2 only
- via neither



3. Interventional effects for MM



- ▶ Vansteelandt and Daniel generalized interventional effects to the multiple mediators setting [Vansteelandt and Daniel, 2016]:

$$\text{RIA-NDE}_1 = E \left\{ \sum_{m_1} \sum_{m_2} E(Y(x, m_1, m_2 | c)) \{P(M_1(x) = m_1 | c) - P(M_1(x^*) = m_1 | c)\} P(M_2(x) = m_2 | c) \right\}$$

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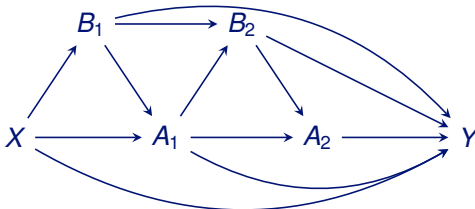
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- ▶ Similar definitions for RIA-NDE and RIA-NIE₂
- ▶ They do not require an ordering of the mediators
- ▶ Identification does not require assumption of no omitted confounders for the mediators
- ▶ **Interpretation:**
 - **RIA-NDE₁**: effect of X on Y via M_1 but not its descendants
 - **RIA-NDE₂**: effect of X on Y via M_2 but not its descendants
 - **Remainder** in the decomposition of TCE, RIA-NIE₁₂, captures the indirect effect of X on the dependence among the mediators

4. Interventional effects for time-varying mediators

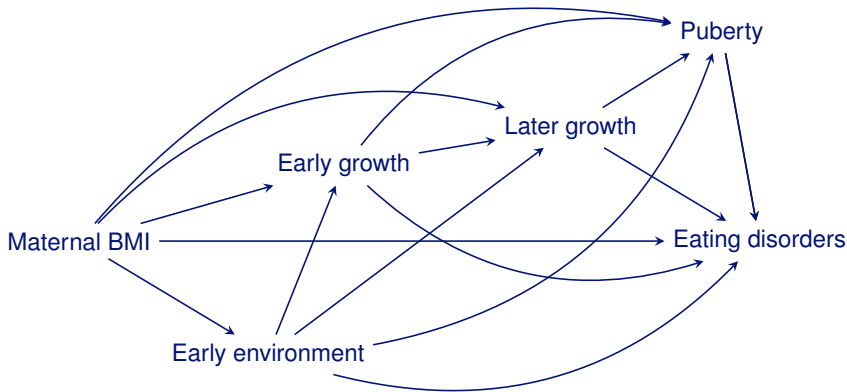
- ▶ Similar extensions of the definition of RIA natural effects given by VanderWeele & Tchetgen-Tchetgen (2017) for the setting with time-varying exposures, mediators, and confounders.
- ▶ A simple example with time-varying mediator A and time-varying confounder B :



- 1 Introduction
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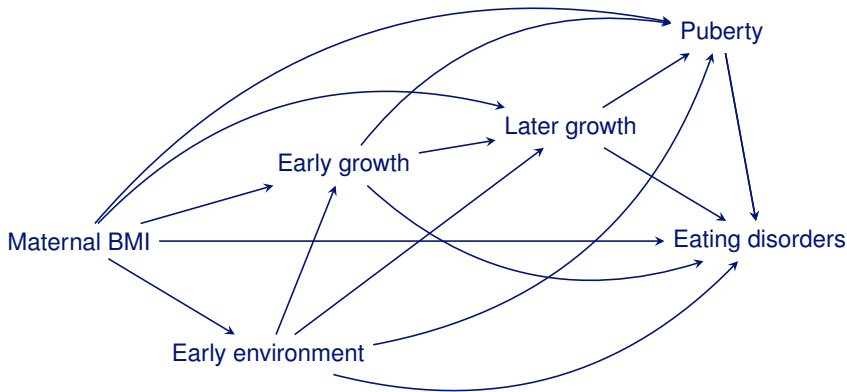
Eating disorders (ED) in adolescent girls

- ED are heterogeneous psychiatric illnesses with typical **onset in adolescence**
- Evidence of **intergenerational dependencies**: maternal BMI
- **Explanations**: environmental factors, childhood metabolic and growth factors



Eating disorders (ED) in adolescent girls

- Q1. Do all mediators combined lead to a null direct effect?
- Q2. Are early factors more important than late factors?
- Q3. Which aspect of growth?

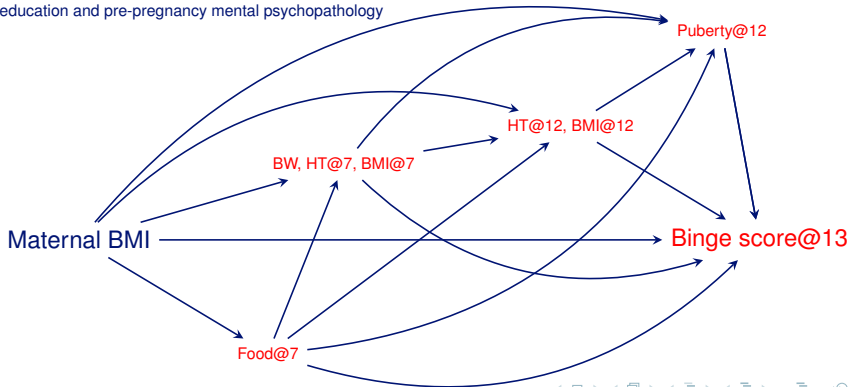


The ALSPAC Study

Cohort of children born in 1991-92 in SW England (3500 girls)



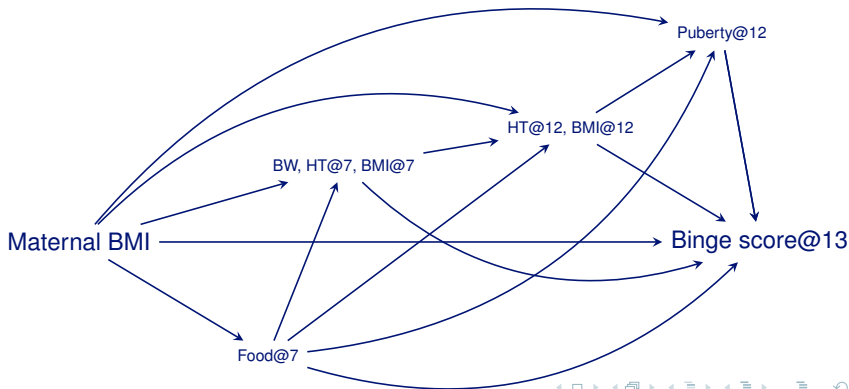
- **Outcome:** *"Binge/overeating"* score, from parental reports @13.5y
- **Exposure:** Pre-pregnancy maternal BMI: BMI <math><25\text{kg/m}^2</math> vs. BMI >math>>25\text{kg/m}^2</math>
- **Mediators:** Birth weight, height and BMI measures, attitude to food, breast development
- **Confounders:** maternal age, social class, education and pre-pregnancy mental psychopathology



1. Natural effects for the 8 mediators jointly

Do all mediators combined lead to a null direct effect?

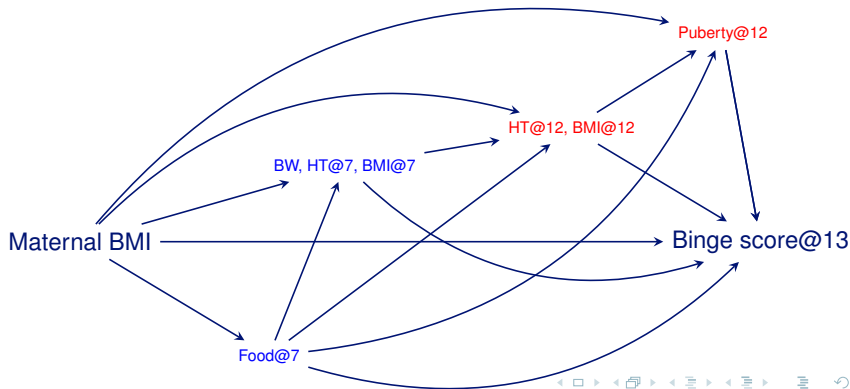
TCE	0.149	(0.079, 0.218)
NDE	0.059	(-0.019, 0.137)
NIE	0.090	(0.026, 0.153)



2. Sequential mediation

Early or late factors?

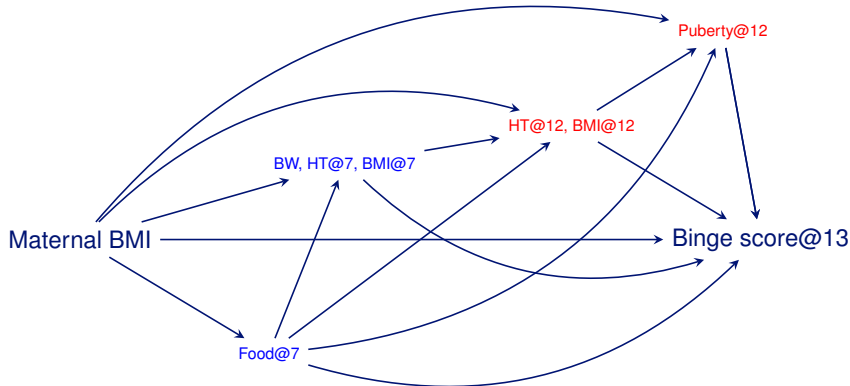
TCE	0.147	(0.081, 0.213)
NIE via M_1 and downstream	0.054	(0.006, 0.102)
NIE via M_2 only	0.034	(0.002, 0.066)



3. Interventional effects

Early or late factors?

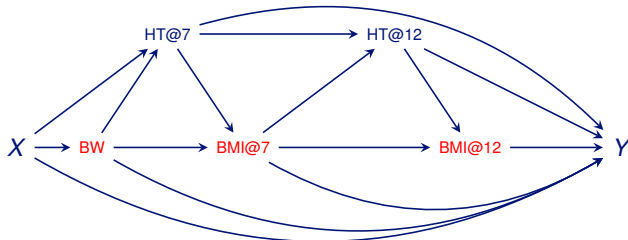
TCE	0.148	(0.096, 0.200)
RIA-NDE	0.059	(-0.001, 0.119)
RIA-NIE via M_1 and upstream	-0.009	(-0.054, 0.035)
RIA-NIE via M_2 and upstream	0.101	(0.050, 0.151)
RIA-NIE via M_1 - M_2 dependence	0.002	(-0.006, 0.009)



4. Interventional effects

Which aspect of growth?

TCE	0.144	(0.091, 0.203)
RIA-NDE	0.063	(-0.008, 0.134)
RIA-NIE via BW and BMI	0.081	(0.035, 0.126)



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Multiple questions



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- ▶ We may not be in a position to answer (all) the questions we wish to ask, especially when unmeasured confounding is suspected
- ▶ However interventional effects offer a useful new perspective as it relaxes the need for ordering and unconfoundedness of the mediators

- Daniel RM, De Stavola BL, Cousens SN, Vansteelandt S. Causal mediation analysis with multiple mediators. *Biometrics* 2015; 71, 1–14.
- Daniel RM, De Stavola BL, and Cousens SN. gformula: Estimating causal effects in the presence of time-varying confounding or mediation using the g-computation formula. *Stata Journal* 2011; 11: 479–517.
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- Vanderweele TJ, Vansteelandt S, Robins JM. Effect decomposition in the presence of an exposure-induced mediator-outcome confounder. *Epidemiology* 2014; 25, 300–306.
- Vanderweele TJ, Tchetgen-Tchetgen E. Mediation Analysis with Time-Varying Exposures and Mediators *JRSS A* (in press)
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